

**The development of gendered speech in children: Insights from adult L1 and L2
perceptions**

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1 **Abstract:** Past studies have shown that boys and girls sound distinct by 4 years old, long before
2 sexual dimorphisms in vocal anatomy develop. These gender differences are thought to be
3 learned within a particular speech community. However, no study has asked whether listeners'
4 sensitivity to gender in child speech is modulated by language experience. This study shows that
5 gendered speech emerges at 2.5 years old, and that L1 listeners outperform L2 listeners in
6 detecting these differences. The findings highlight the role of language-specific sociolinguistic
7 factors in both speech perception and production, and show that gendered speech emerges earlier
8 than previously suggested.

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10 *Keywords:* speech perception, children's voices, gender identification, language experience,
11 acoustic information

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12 **1. Introduction**

13 The voices of adult male and female humans are typically very easy to distinguish. For
14 example, the average f_0 for adult males is about half that of the average f_0 for adult females (e.g.,
15 Baken & Orlikoff, 2000). This f_0 difference, along with other salient acoustic differences such as
16 lower spectral frequencies in males than females, is directly linked to anatomical traits: males tend
17 to have long vocal tracts and more slowly vibrating vocal cords than women. But interestingly,
18 past studies have shown that listeners can classify voices of children as young as 4-year-olds as
19 belonging to a boy or girl – long before anatomical differences between the sexes emerge, and that
20 these decisions are at least in part related to acoustic parameters such as f_0 and formant frequencies
21 (e.g., Perry et al., 2001). The most plausible explanation for why boys and girls sound distinct
22 before adolescence is their early adoption of community-specific gender differences in the way
23 children talk. Indeed, sociolinguistic studies have shown that sensitivity to gender differences in
24 speaking style is early emerging (e.g., Ladegaard & Bless, 2003). In the current study, we use a
25 set of longitudinal recordings to examine 1) when perceptible differences between boys' and girls'
26 productions first emerge, and 2) whether listeners' sensitivity to the cues distinguishing boys and
27 girls is affected by language experience.

28 Because gender differences in speech are at least in part learned from the speech
29 community (i.e., not only due to anatomical differences; Johnson, 2005), the cues differentiating
30 prototypical male versus female speech can differ across dialects and languages. For example, f_0
31 differences between males and females have been shown to be more pronounced in Japanese
32 speakers than in Dutch speakers (van Bezooijen, 1995). And Korean-English bilinguals change
33 their f_0 range to match the speech patterns of the language they are speaking, with males using a
34 wider f_0 range when speaking Korean, but a narrower range when speaking English (Cheng,

35 2020). Beyond f_0 , language-specific gender differences have been observed in the acoustic-
36 phonetic realization of speech segments. For example, gender differences in the production of /s/
37 have been reported to be more pronounced in Japanese than in English (Heffernan, 2004). And
38 gender-based differences in VOT have also been argued to vary by language. For example,
39 female Seoul Korean speakers reportedly produce shorter VOTs for aspirated stops than males
40 (Oh, 2011), whereas female English speakers reportedly produce longer VOTs than males (Robb
41 et al., 2005; see, however, Morris et al., 2008).

42 Taken together, the examples given above demonstrate that gendered speech patterns are
43 not solely physiological in nature, but are also learned through socialization within a particular
44 community. Interestingly, evidence from the developmental literature has shown that children
45 adopt these gender differences in speaking style early in life. In a cross-sectional study, native
46 English adult listeners who were presented with recordings from 4-, 8-, 12-, and 16-year-olds,
47 performed above chance at identifying gender in even the youngest of these children (Perry et
48 al., 2001), despite the 4- and 8-year-olds being too young to exhibit sex differences in vocal tract
49 anatomy (Tecumseh Fitch & Giedd, 1999). And some studies have reported gender differences
50 in formant frequency values before puberty, arguably because boys attempt to produce a more
51 masculine voice and speech pattern by lowering their jaw and modifying the extent of lip
52 rounding (e.g., Lindblom & Sundberg, 1971; Bennett & Weinberg, 1979; Bennett, 1981). Other
53 studies have suggested that gender differences in the realization of /s/ are present in children as
54 young as 4 years of age (Li et al., 2016). And Yang and Mu (1989) reported acoustic differences
55 in the speech of three years old, the youngest reported differences we are aware of. To
56 summarize, perception studies suggest that boys and girls produce perceptible differences in their

57 speech by the age of four, and production studies suggest that gender differences in speech may
58 emerge even earlier – perhaps by the age of 3.

59 Despite well-documented differences in production between male and female speech, no
60 study to date has asked whether these translate into language-specific differences in the
61 perception of speaker gender. In other words, no study has asked whether listeners are more
62 accurate at identifying speaker gender in speakers of their own language community than
63 speakers from another language community. Perhaps this is because physiological cues to
64 speaker sex (e.g., average f_0) are so salient in adult speech that they might overwhelm any
65 differences due to language-specific experience. But as outlined above, gendered speech patterns
66 in prepubertal children are thought to be learned (i.e., not due to physiological differences
67 between boys and girls). Thus, one question that is ripe for investigation is whether a listeners'
68 ability to distinguish boys and girls depends on the listener's language background.

69 In the current study, we ask when perceptible differences in boys' and girls' productions
70 emerge, and if sensitivity to gender differences in children's speech is tied to a listener's
71 language experience. To accomplish this, we use a set of recordings collected longitudinally,
72 with the same set of native English-speaking children recorded at 2.5, 4, and 5.5 years of age.
73 We then tested L1 English and L2 English speakers' accuracy at identifying these children's
74 gender at these three ages. We hypothesize that adults should perform above chance at
75 classifying speech productions as belonging to boys and girls by at least 4 years of age, and that
76 performance in differentiating boys from girls should improve with child age. In addition, we
77 hypothesize that language experience should modulate listeners' ability to classify children's
78 gender. Thus, we predict L1 listeners will be more accurate than L2 listeners in both classifying
79 child gender and gauging their performance (i.e., confidence rating more in line with accuracy).

80 2. Method

81 2.1 Participants

82 Forty-eight adults from the University of Toronto participated in the perception study.
83 Half were L1 learners of English (4 males, $M_{age} = 20.8$) and half were L2 learners of English (12
84 males, $M_{age} = 19.52$). All L1 learners acquired English in Canada before the age of 6, and used
85 English at least 80% of the time; all L2 learners moved to Canada after the age of 14 and had
86 minimal to no classroom exposure to English before the age of 14. L2 learners' first languages
87 were Mandarin (18) and Cantonese (6). Participants reported no hearing or vision impairments at
88 the time of testing.

89 2.2 Stimuli

90 Stimuli were drawn from a child speech corpus consisting of isolated words, elicited
91 using an experimenter-controlled video game (Cooper et al., 2020). Children viewed a target
92 word on a computer screen (e.g., ball, duck) and were prompted to name the picture in citation
93 form. Twenty-four words were chosen for use in the current study. All chosen words were
94 produced by the same 12 Canadian English-learning children (6 assigned male at birth and 6
95 assigned female at birth) at each of three different ages: 2.5, 4, and 5.5 years old. See
96 Supplementary Material² for a full set of experimental words. These children were identified as
97 cisgender by their primary caregiver(s). Stimuli were normalized for root mean square amplitude
98 in Praat 6.0.22 (Boersma & Weenink, 2020).

99 2.3 Procedure

100 Listeners were tested individually in a quiet room using PsychoPy3 (Peirce & MacAskill,
101 2018). On each trial, a word was presented once, and listeners were asked to indicate the child
102 speaker's gender by clicking on a male or female icon (similar to the icons used to indicate

103 bathrooms, and additionally marked by the colours blue and pink)³. To assess whether language
104 experience affected listeners' confidence in their performance, participants were also asked to
105 rate their confidence on a scale from 1 (not at all) to 5 (very confident).

106 To ensure that participants understood the task, the experimental trials were preceded by
107 two practice trials (consisting of two children voices that were not included in the experimental
108 trials). Across the experiment, participants heard 24 words with each word produced by a
109 different boy/ girl pair at 2.5, 4, and 5.5 years of age. Thus, each listener heard tokens from all 12
110 children, each at three ages over the course of 144 trials (24 Words \times 2 Gender \times 3 Ages). The
111 trial order was randomized for all participants. The study took approximately 10 minutes to
112 complete.

113 **3. Results**

114 *3.1 Perception*

115 To assess L1 and L2 listeners' performances across the three children's ages, we fit a
116 generalized mixed-effects model to our data using the *glmer* function of the *lme4* package (Bates
117 et al., 2015) in R. The binary response variable was Accuracy (1 = correct response). The
118 independent variables, listener Group and children's Age, were entered as fixed effects. An
119 interaction term between the two fixed effects was not included because, for the purpose of the
120 study, we are primarily focused on the difference between L1 and L2 listeners' abilities (not the
121 relative magnitude of this difference across age). We included random intercepts for Word, and a
122 random slope for children's Age by Participant. Listener Group was simple-coded (with L1
123 listeners as the reference level). In addition, because we expected listeners would be more
124 accurate with older children than with younger children, we coded children's Age with Helmert
125 contrasts: 1) 2.5-year-olds vs. 4- and 5.5-year-olds combined, and 2) 4-year-olds vs. 5.5-year-

126 olds. The β -coefficient for the intercept represents the log odds of a correct response averaged
 127 across all ages, and the β -coefficient corresponding to each effect represents the difference in log
 128 odds of a correct response between the two levels of that comparison, collapsed over all levels of
 129 the other factor.

130 As predicted, a significant intercept was found, indicating listeners' overall performance
 131 above chance (see Table 1 and Fig. 1a). Moreover, the model revealed a significant main effect
 132 of listener Group, with L1 listeners performing better than L2 listeners (irrespective of children's
 133 age)⁴. The model also revealed that listeners' performance differed significantly between 2.5-
 134 year-olds vs. older children, but no difference was found between 4- and 5.5-year-olds.
 135 Importantly, a subsequent follow-up test, using the same model but with 2.5-year-olds coded as
 136 the reference level, shows that listeners' performances with 2.5-year-olds were significantly
 137 above chance ($\beta = 0.57$, $SE = 0.09$, $z = 6.67$, $p < 0.001$). Note that this is the first study to date to
 138 demonstrate that children this young already produce perceptible gender differences in their
 139 speech.

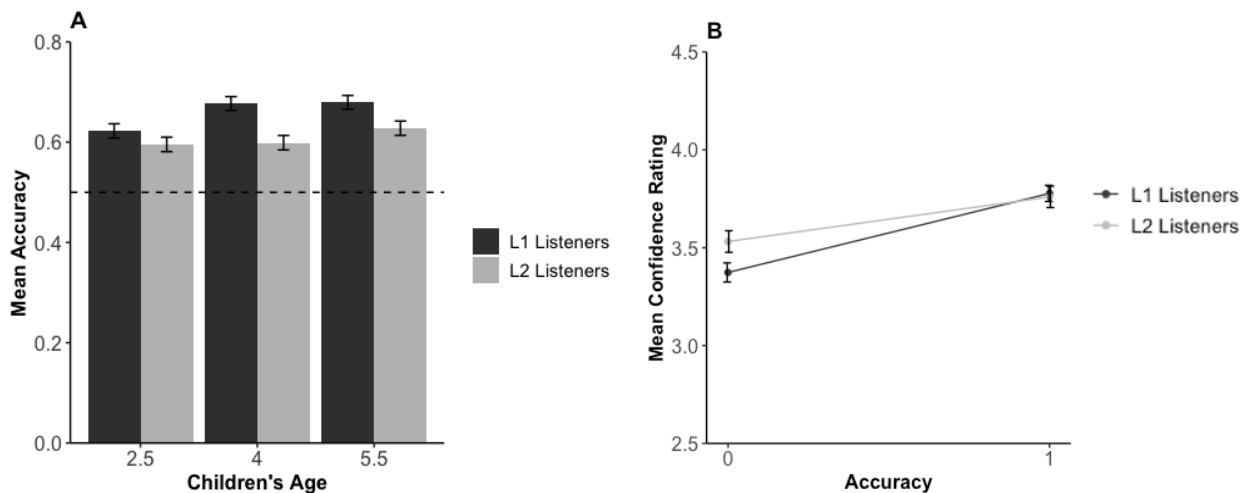
140 Table 1. Summary of results from a logistic regression model for gender classification.

	β	SE	z	p
Intercept	0.68	0.08	8.67	< 0.001
Group: L2 listeners (vs. <i>L1 listeners</i>)	-0.22	0.05	-4.11	< 0.001
Age 2.5 (vs. <i>Age 4 & Age 5.5</i>)	0.16	0.06	2.88	0.004
Age 4 (vs. <i>Age 5.5</i>)	0.07	0.07	1.01	0.31

141 *Note.* Model: `glmer(accuracy ~ group + age + (1 | word) + (age | participant))`. Italics indicates the reference level.

142 In addition, we analyzed whether listeners were able to gauge their performance
 143 accurately. To assess this, we used a linear mixed-effects model to predict Confidence rating
 144 from Accuracy, listener Group, and their interaction. Listener Group was simple coded (with L1

145 listeners as the reference level) and Confidence rating was centred. The model showed that both
 146 Accuracy ($\beta = 0.39$, $SE = 0.04$, $t = 9.91$, $p < 0.001$) and listener Group ($\beta = 0.17$, $SE = 0.04$, $t =$
 147 3.90 , $p < 0.001$) significantly predicted Confidence rating. A significant interaction was also
 148 found ($\beta = -0.18$, $SE = 0.06$, $t = -3.33$, $p < 0.001$), with L1 listeners ($M = 3.37$, $SE = 0.05$)
 149 reported lower confidence than L2 listeners ($M = 3.53$, $SE = 0.05$) for incorrect responses, but
 150 equally high confidence for correct responses (L1 listeners: $M = 3.78$, $SE = 0.04$; L2 listeners: M
 151 $= 3.77$, $SE = 0.05$)⁵. The results indicate that, overall, listeners were able to gauge their
 152 performance accurately, but the ability to gauge their performance when incorrect was
 153 modulated by language experience (see Fig. 1b).



154
 155 Fig. 1. a: Mean accuracy of gender classification of 2.5-, 4-, and 5.5-year-olds by L1 and L2 listeners. Dashed line at
 156 0.5 indicates the chance level. b: Mean confidence rating of L1 and L2 listeners when responses were correct
 157 (Accuracy = 1) and incorrect (Accuracy = 0). Error bars indicate SE based on by-participant means.

158 3.2 Acoustics

159 We conducted follow-up tests to examine 1) whether children in our sample exhibited
 160 gender and age differences on several acoustic measures (f_0 , F_1 , F_2), and 2) which, if any, of
 161 these acoustic measures are predictive of L1 and L2 listeners' responses, focusing on
 162 monosyllabic words to reduce variability. We manually annotated the vocalic portion of all

163 tokens (including the vowel as well as preceding/following /l/ or /r/ when present). Given the
164 inherent difficulty in estimating f0 and formant values in children's speech using static
165 parameters, we manually inspected each token and chose the optimal parameters for obtaining
166 accurate acoustic measures for f0 and formants for each token independently. Using these
167 parameters, measures of f0, F1, and F2 were taken at the midpoint of the vocalic interval. Five
168 words with diphthongs or coda /r/ were excluded from formant analyses due to the dynamic
169 nature of the formants. Tokens for which was not possible to obtain an accurate pitch (8) and/or
170 formant track (27) were omitted from the relevant analyses, and mispronounced tokens (2) were
171 omitted entirely.

172 To examine whether acoustic measures differed by gender and age, we performed
173 separate analyses for f0, F1, and F2 (see Fig. 2) using linear mixed-effects models (implemented
174 using the *lmer* function of the *lme4* package (Bates et al., 2015) whereas *p*-values were computed
175 using the *lmerTest* package; Kuznetsova et al., 2017). In each model⁶, the acoustic dimension
176 was the response variable, with Gender, Age, and their interaction as fixed effects. Both
177 predictors were simple-coded, with reference levels as Male for Gender and 4-year-olds for Age.
178 β -coefficients represent mean differences in the acoustic value between levels of the predictor
179 factor. Only significant and trending main effects or interactions ($p < .010$) are reported here. As
180 expected, the model for f0 revealed a significant decrease across the three ages ($\beta = 18.79$, $SE =$
181 9.89 , $z = 1.90$, $p = 0.03$). Although there was no significant effect of Gender, a significant Age \times
182 Gender interaction was found in the comparison of 4- and 5.5-year-olds ($\beta = -45.10$, $SE = 18.58$,
183 $z = -2.43$, $p = 0.02$). Wald χ^2 test was performed using the *phia* package to explore this
184 significant interaction (De Rosario-Martinez, 2015). A follow-up test indicated that 5.5-year-
185 olds, but not 4-year-olds, showed a marginal gender difference (5.5-year-olds: $\chi^2 = -23.49$, $p =$

186 0.056; 4-year-olds: $\chi^2 = -5.21, p = 0.37$). For F1, a marginally significant difference in the
187 expected direction was found between 4- and 5.5-year-olds ($\beta = -72.54, SE = 32.66, z = -2.22, p$
188 $= 0.07$). In summary, we found only marginal gender differences, and only in 5.5-year-olds, with
189 both in the expected direction. However, it is important to note that because stimuli were chosen
190 for ease of production by 2.5-year-olds, these analyses were therefore based on a relatively small
191 amount of data; given the low power, these null results should be interpreted with caution.

192 Next, we examined how well the acoustic measures discussed above predicted L1 and L2
193 listeners' responses. We performed separate analyses for f0, F1, and F2 using a generalized
194 mixed-effects model⁷ in R. The model predicted listeners' Responses of "female" (vs. "male")
195 from the acoustic measures (f0/F1/F2) and listener Group as well as their interaction. Listener
196 Group was simple coded (with L1 listeners as the reference level). As above, we included
197 random intercepts for Word and Participant. The β -coefficient represents the difference in log
198 odds of a "female" response between the two levels of that comparison, collapsed over all levels
199 of the other factor. The model revealed that, overall, listeners' responses are predicted by f0 ($\beta =$
200 $0.01, SE = 0.001, z = -7.86, p < 0.001$), F1 ($\beta = 0.003, SE = 0.0003, z = 8.23, p < 0.001$), and F2
201 ($\beta = 0.001, SE = 0.0001, z = 6.88, p < 0.001$), with low values eliciting mostly "male" response
202 and high values eliciting mostly "female" response. This is consistent with the literature that
203 lower pitch and formants are associated with males' voices. However, no difference was found
204 between L1 and L2 listeners for any of the three dimensions (i.e., there was no significant
205 interaction between Group and any of the acoustic measures).

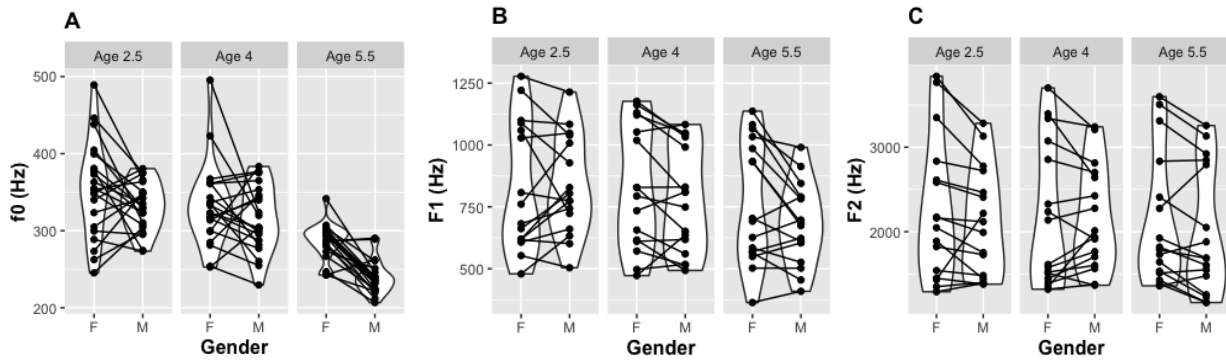


Fig 2. a-c: Gender differences in f_0 , F_1 , and F_2 values across the three ages in children.

4. Discussion

Past work has demonstrated that gender differences in children's speech are perceptible by 4 years of age, long before anatomical differences in children's vocal tract anatomy are generally thought to have developed (see Munson & Babel, 2019, for discussion). These gender differences in speech are thought to be learned early in life, with what constitutes typical male or female speech patterns differing to some degree across language communities. In the current study, we investigate the possibility that gendered speech patterns can be detected in children even younger than 4 years of age, and also ask whether adult listeners' sensitivity to these patterns is dependent on language experience.

In line with our predictions, we found that L1 listeners outperformed L2 listeners in their gender classification of child voices. Moreover, both L1 and L2 listeners performed well above chance in classifying the gender at all three age groups tested, including the 2.5-year-olds. As far as we are aware, this is the first published study to demonstrate an apparent role of language experience in gender classification, and also the first study to show that children as young as 2.5 years of age show perceptible gender differences in their speech productions. Our findings with L1 listeners fit well with other longitudinal studies on the development of gendered speech patterns (Munson et al., 2019). In our follow-up acoustic analyses, we also found that all three

225 acoustic measures (f_0 , F1, F2) influenced listeners' judgments of children's gender, although we
226 found only marginal evidence of gender-based acoustic differences. This suggests that the task
227 might have tapped into participants' stereotypes about gender differences in adults, rather than to
228 their knowledge of differences between boys' and girls' speech.

229 Given the weak relationship between our acoustic measures and child gender, how did
230 listeners perform above chance in our gender classification task? One possibility is that listeners
231 based their classifications on aspects of the signal that we did not measure. This would be
232 consistent with claims that early gender classifications might be based on more holistic qualities
233 such as voice quality. For example, Günzburger and Keurs (1987) found that listeners judged
234 girls' voices to be significantly clearer, softer, shriller, high-pitched, melodious and precise than
235 boys' voices. In contrast, these same listeners described boys' voices as duller, louder, deeper,
236 more monotonous and more careless than girls' voices. These descriptions are in line with a
237 rating study we are currently running with these same recordings, where participants tend to
238 describe boys' and girls' voices with similarly stereotyped labels. It is possible that these labels
239 reflect community norms for gendered speaking styles, and that they could have driven listeners'
240 classification decisions in our study. Additional support for this view comes from previous
241 studies also failing to find robust spectral differences in children under 6 years of age (e.g., Lee
242 et al., 1999; Perry et al., 2001; Vorperian et al., 2019).

243 Our finding that listeners' language experience impacts how accurately they can classify
244 boys' and girls' voices fits well with the hypothesis that gendered speech in children is learned
245 (i.e., not simply due to biological maturation). This is also consistent with the view from the
246 linguistic literature that variations in speech patterns reflect the process of social differentiation
247 (Eckert, 2012; Foulkes et al., 2010). Interestingly, we found that L1 and L2 listeners both relied

248 on f_0 , F1, and F2 when making gender judgments. Therefore, other characteristics of the speech
249 signal that are not measured here may be responsible for this performance differential between
250 L1 and L2 listeners. A likely explanation, given the literature on the language-specific nature of
251 gender differences in speech, is that the L1 listeners had knowledge of language-specific cues to
252 talker gender that the L2 listeners lacked. Presumably, these cues were not captured in our
253 acoustic measures. Future studies could address this issue more directly by manipulating specific
254 cues to talker gender that differ between two language populations (e.g., see if Japanese speakers
255 are more likely to classify children with a wider f_0 as boys whereas English speakers are more
256 likely to classify them as girls), as opposed to previous perception studies which only focused on
257 English speakers. However, another possible explanation for performance differences between
258 L1 and L2 listeners is that the latter were distracted by attempts to understand the children in this
259 study, giving them few processing resources to devote to gender classification. We find this
260 explanation unlikely, however, because our task did not require listeners to comprehend child
261 talkers. Nonetheless, future studies could test this possibility by having the same target word or
262 phrase used on every gender classification trial.

263 To conclude, gender differences in children's speech clearly emerge even earlier than
264 previous work has demonstrated, but L1 listeners are better at detecting these differences than L2
265 listeners. Logically speaking, this may be because at least some of the realization of gender in
266 children's speech is language-specific. These findings generate many hypotheses about the
267 sociolinguistics of early speech acquisition, and make exciting predictions regarding
268 communication and development in linguistically-diverse communities.

269 **Acknowledgements**

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 271 Humanities Research Council of Canada and the Natural Sciences and Engineering Research
 272 Council of Canada.

273 **Footnotes**

274 ¹ Note that the cited literature on this topic uses the terms gender and sex loosely, often failing to
 275 include a robust measure of gender identity, and not give an explanation for how talker sex was
 276 determined.

277 ² See supplementary material at [URL will be inserted by AIP] for list of experimental words.

278 ³ We are aware that forcing participants to make a binary decision does not reflect the complex
 279 realities of gender identities in the real world, but this binary classification task was appropriate
 280 for addressing our question of interest in the current study.

281 ⁴ Although it was not included in the initial analysis, based on a reviewer's comment, we tested
 282 whether there might be a children's Age \times listener Group interaction. We found a trending
 283 interaction ($p = 0.07$). Upon breaking down the interaction, we found that the effect was
 284 significant for 4- and 5.5-year-olds but not for 2.5-year-olds. Although we cannot draw firm
 285 conclusions given that the interaction was only trending, this suggests that looking into the
 286 trajectory of this effect across ages is an interesting topic for future work.

287 ⁵ The same main results were found when the data was modeled using a Poisson distribution.

288 ⁶ Model: $\text{lmer}(f0/F1/F2 \sim \text{gender} \times \text{age} + (1|\text{word}) + (1|\text{participant}))$

289 ⁷ Model: $\text{glmer}(\text{female.response} \sim f0/F1/F2 \times \text{group} + (1|\text{word}) + (1|\text{participant}))$

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